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# Corrective Adjustment Procedures as a strategy to remove Relative Age Effects: Validation across male and female age-group long jumping

This is a pre print version of the following article:			
Original Citation:			
Availability:			
This version is available http://hdl.handle.net/2318/1869846 since 2023-08-02T08:43:32Z			
Published version:			
DOI:10.1016/j.jsams.2022.04.007			
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## Abstract

1

2	Objectives: To validate the application of Corrective Adjustment Procedures (CAPs) to remove
3	Relative Age Effects (RAEs) in youth athletic contexts, this study estimated the longitudinal
4	relationships between decimal age (chronological and relative) and performance in male and female
5	long jumpers. Using trendlines, CAPs were applied, and RAE distributions associated with performance
6	attainment were re-examined.
7	Design: Retrospective longitudinal design examining publicly available long-jump competition
8	performance data between 2005-2019.
9	Methods: In part I, participants were 689 junior Italian long jumpers (age range=11.01-17.99 years;
10	56.6% females) who participated in $\geq$ three events. Longitudinal modelling and regression equations
11	quantified the sex-specific relationships between decimal age and long jump performance. In part II,
12	equations were utilised to adjust individual performance within an independent sample (N=13,639;
13	50.1% females) of age-matched jumpers. RAE distributions within attainment levels (i.e., Top 25-10%)
14	were examined based on raw and correctively adjusted performance.
15	Results: Irrespective of sex, RAEs were prevalent across all age-groups with medium-large effect sizes
16	at 12-16 (males) and 12-14 years (females) of age (Cramer's V range = 0.19-0.34). RAE bias magnitude
17	also increased with attainment level (i.e., Top 25-10%). Following CAPs application, typical RAEs
18	were removed with non-significant deviations in relative age distributions regardless of sex, age-group
19	or attainment level (Top 25 or 10%).
20	Conclusions: Based on sex-specific longitudinal reference data, findings provide efficacy for CAPs
21	application to remove RAEs across youth long jumping events. CAPs suggest potential in improving
22	performance evaluation, identification of technically skilled performers, and general sporting
23	experiences.
24	

Keywords: Relative Age Effect, Corrective Adjustment Procedures, Track & Field, Athletics, Athlete
Development.

#### 27 Introduction

28 With the purpose to efficiently organise participation opportunities, structure youth sport competition, and actually reduce inter-individual developmental differences, sport governing 29 30 bodies/federations commonly group athletes according into (bi)annual age-group cohorts.<sup>1-3</sup> Unfortunately however, (bi)annual age-grouping still permits the possibility of chronological age 31 32 associated differences of up to 12 (or 24) months between individuals, often leading to Relative Age 33 Effects (RAEs).<sup>4-6</sup> Spanning from initial 'grassroots' to adult professional levels, RAEs reflect a highly prevalent participation and attainment inequality within youth sport and educational settings.<sup>5-7</sup> 34 Specifically, RAEs are characterised by asymmetry, where relatively older individuals within a given 35 age-group cohort are overrepresented, while the relatively younger are underrepresented.<sup>5,8</sup> Although 36 37 causes are still somewhat being verified, likely accountable factors (to a greater or lesser extent depending on context) include: normative age-based differences in anthropometric development (e.g., 38 height & weight);5,7 potentially added biological maturational variation during maturational years;1 39 parallel neurological and cognitive developmental differences;<sup>5,6</sup> social evaluation processes (e.g., coach 40 evaluation relative to age-group peers)9 and their consequences upon individual psychology (e.g., 41 perceived competence, value).<sup>6,10</sup> Within youth sport, competition tiers or talent identification practices 42 also encourage selective differentiation, and as a consequence, RAEs can be further magnified<sup>11,12</sup> with 43 the relatively older consistently outperforming their counterparts in youth rankings13,14 or being 44 45 significantly more likely to be selected. Such trends are particularly evident in sport contexts with high physiological demands, such as team (soccer, rugby, basketball)<sup>8,15,16</sup> and individual sports (e.g., track 46 and field).1,17 47

Related to track and field, several international studies have identified significant over- and under-representations of the relatively older and younger respectively across male and females at different age groups and differing levels of competition.<sup>1,17-20</sup> Here, RAEs effect sizes have been identified as being larger in adolescent age ranges (e.g., Q1 40% vs. Q4 13%), than post-18 age-groups (e.g., Q1 39% vs. Q4 17%) or adult competition (e.g., Q1 35% vs. Q4 19%)<sup>18</sup> and were further magnified according to selection level.<sup>18,21</sup> Moreover, RAEs are generally larger in males relative to females in various sport contexts.<sup>7</sup> An extensive cross-sectional study of international track and field did identify that males had larger RAE magnitude both in junior and senior categories.<sup>17</sup> However, the greatest RAE magnitude in females has been associated with preadolescent (less than 11 years old) and adolescent (12–14 years old) age categories, with declining risk following maturation.<sup>7,22</sup> This differential timing of peak RAE magnitude has been attributed to earlier chronological age time points for (re-)accelerated growth, less heightened anthropometric and physiological inter-individual differences during maturational years, and, earlier age time points for attaining adult stature relative to males.<sup>7</sup>

61 To address RAE participation and attainment inequalities, a range of feasible strategies have been proposed.<sup>23-26</sup> Corrective Adjustment Procedures (CAPs) appears to be a promising strategy, 62 particularly for individual sports where performance is quantified, such as track and field<sup>1,21</sup> or 63 swimming.<sup>24,27</sup> During their examination of 8-15 year-old sprinters, Romann and Cobley<sup>1</sup> originally 64 65 applied CAPs following identification of consistent RAEs. Using a regression trendline summarising the decimal age-performance relationship, they quantitatively adjusted individual performance times 66 given the expected time for the relatively oldest individual in an annual cohort. Following CAPs 67 application, the distributions of who attained the corresponding Top 50-10% of performance times were 68 re-examined. They identified RAEs as being completely absent in the Top 10% sprinters and were either 69 70 removed or reduced in the Top 50%-25% of sprinters. Since, similar trial results have been obtained when examining youth world-class sprinters,<sup>21</sup> Australian male 100 m freestyle<sup>24</sup> and female 100 and 71 72 200 m breaststroke swimmers.<sup>27</sup> Here, CAPs were able to predominantly remove moderate-large RAEs 73 across age-groups and artificially-created selection levels. These findings provide some initial efficacy 74 for CAPs as a mitigation strategy against RAEs<sup>24,27</sup> and help more accurately evaluate (alternative) factors influencing youth age-group athletic performance.<sup>21</sup> 75

While prior studies initially developed or tested CAPs in sprinters and swimmers, their validity and potential in other contexts still requires evaluation. Thus, the purposes of the present study (Part I) were to estimate the longitudinal relationships between decimal age (i.e., chronological, and relative) and performance across youth annual age-groups female and male long jumpers. Then, to determine whether CAPs could remove RAE performance attainment inequalities across youth annual age-groups female and male long jumpers (Part II).

82

#### 83 Methods

#### 84 Initial data extraction

All data were collected from the FIDAL (Italian Track and Field Federation; 85 http://www.fidal.it/) database and included long jump performance for males and females at officially 86 recognised competitions (i.e., where official measurements were recorded). Aligning with FIDAL 87 88 classification, long jumpers competing across groups 12-17 ages inclusive were included. For each age 89 category, long jumpers ranked in the Top 150 official lists been 2005-2019 (inclusive) were available for data extraction (principal database). The data reported individual seasonal best performance, name, 90 day of birth, club, competition date, competition venue, and competition date. Only results obtained 91 with legal wind speed (≤ 2 m/s) were included aligning with World Athletics rules.<sup>28</sup> This study was 92 93 approved by the local ethics committee of the University of Torino and conducted according to the declaration of Helsinki. 94

## 95 Part I

96 To determine the relationship between decimal age and performance, a subset of participants with longitudinal data were extracted and examined from the principal database. Participants were 299 97 98 male and 390 female long jumpers, who at the time of competition were 11.01-17.99 years of age, and 99 who recorded a valid performance (i.e., seasonal best) in at least three years (not necessarily 100 consecutive). Across longitudinal tracking, the male sample contributed 1,326 performances, while 101 females contributed 1,740 performance distances. The examination of longitudinal data was done to 102 more accurately estimate jump performance progressions over time (i.e., within person change).<sup>10,24,27</sup> 103 Screening for data entry errors and outliers was conducted, with outliers (i.e., Z-score values > -2) of low performance removed prior to analysis. The exact age, based on the year and day of athletes' 104 birthdate, at which athletes achieved their performances distances was calculated. 105

To determine sex-specific decimal age-performance relationships, the exact decimal age (i.e., year and date of competition; independent variable) and longest jump performance (m; dependent variable) were examined independently for males and females using mixed model regression. The best fitting model trendline (i.e., linear [y = ax + c] vs quadratic  $[y = ax^2 + bx + c]$ ) was evaluated. Decimal age was entered as a fixed factor, while participants were entered as a random factor. Model fit was assessed using the likelihood ratio test. The best fitting model and corresponding regression equation

112 was subsequently taken forward for CAPs calculations used in Part II.

113 Part II

To determine whether corrective adjustments could remove RAEs, we re-examined the original database which included jumpers (N = 6,812 males and N = 6,827 females, aged 12–17 years) who were ranked in the Top 150 official lists of Italian National federation (FIDAL) from 2005-2019. Participants' age, date of birth, and individual seasonal best performance were again extracted from the database.

To determine whether RAEs existed in raw performance distributions, all long jumpers were 118 categorized according to annual-age group and relative age quartile. With reference to FIDAL cut-off 119 date criteria for age-grouping, the following dates were utilised: the relatively oldest were participants 120 121 born in January-March = Quartile 1 (Q1); those between April-June = Quartile 2 (Q2); July-September = Quartile 3 (Q3); while the relatively youngest were those born between October-December = Quartile 122 4 (Q4). To identify RAEs, differences between observed and expected uniform quartile distributions 123 (i.e., 25% for each quartile) were assessed using Chi-square ( $\chi^2$ ) tests, with p set at < 0.05, and effect 124 magnitudes determined by Cramer's V. The threshold values for effect size statistics were:  $0.06 \le V$  for 125 a trivial effect;  $0.06 < V \le 0.17$  for a small effect; 0.17 < V < 0.29 for a medium effect; and  $V \ge 0.29$  for 126 a large effect.<sup>29</sup> Odds Ratios (ORs) and 95% Confidence Intervals [95% CIs] then identified 127 128 discrepancies between Q1 v Q4 and for Q1+2 v Q3+4 (i.e., half-year distribution comparisons). To 129 assess whether RAE effect sizes changed according to attainment level within each age group, RAEs was calculated according to the Top 25% and 10% of jump performances respectively. 130

131 To determine whether RAEs existed in correctively adjusted performance distributions, all raw performances were firstly adjusted using the sex-specific reference equations from Part I. Specifically, 132 individual long jump performance at given decimal age were adjusted based on the mean expected 133 134 performance difference per day. Following this procedure all individual performances were centred on the exact age of a given age-group. For example, considering the 12 years age-group, an athlete can 135 record their seasonal best performance in a decimal age ranged between 11.01-12.99 years. Thus, the 136 137 adjusted performance corresponds to the raw performance plus/less the expected performance 138 differences per day. Specifically, individual long jump performances recorded between 11.01 and 11.99 years old were increased, while individual long jump performances, recorded between 12.01 and 12.99 years old were decreased. Differently, performances recorded when athlete was 12.00 years old were not changed. Then, following adjustments, the relative age quartiles distributions (i.e., Q1-4) were reexamined using similar analytical procedures (as described above). Again, sex-specific distributions across age-groups and according to attainment level were assessed, permitting comparisons with raw performance distributions. All analytical steps were performed using custom-written software in MATLAB R2021a (Mathworks, Natick, MA, USA).

146

147 Results

148 Part I

149 Compared to a linear trendline, a quadratic model fit was identified as significantly better-fitting 150 in summarising the decimal age – performance relationship ( $\chi^2 < 0.05$ ). The variance explained by fitted 151 models was 0.92 and 0.87 for male and female jumpers, respectively. Figure 1 summarises the quadratic 152 trendlines between decimal age (i.e., year and day) and performance for males (Figure 1a) and female 153 jumpers (Figure 1b).

The estimated performance change (m) and percentage change according to age-group (i.e., 12-17 years) is summarised in Table 1. The estimated performance difference within a given age-group decreased consistently with age (i.e., males = 47.51% at 12 to 5.37% at 17 years old; females = 28.12% at 12 years to 0.84% at 17 years of age).

<Insert Figure 1 about here>

159

154

<Insert Table 1 about here>

- 160 Part II
- 161

#### <Insert Table 2 about here>

Table 2 summarises both the raw performance and correctively adjusted relative age quartile distributions, including Chi-square statistics, Odds Ratio from 12-17 years of age for male and female jumpers. In Italian athletics, as there is a bi-annual age grouping process, the prevalence in the official <u>lists of athletes 13 and 15 years old is greater than those of 12 and 14 years old (see Table 2).</u> Findings identified that RAEs were evident across raw male and female samples, with notable prevalence at 12-

#### Commentato [GB1]: Please check

**Commentato [SC2R1]:** Gennaro, I would try to avoid mention or talking about the bi-annual grouping process in the paper – unless mentioned elsewhere and explained well. Will introduce confusion and further comments from reviewers – possibly!?

Which item or issue are you responding too – I couldn't match it on the reviewer comments.

If you address the concern without mention of the dual-year grouping, I would do that.

167 16 years of age with medium-large effect size in males. At similar age-groups, small-medium RAE 168 effect sizes were apparent in females. Quartile comparison odds ratios decreased with age in both sexes 169 (see Table 2 - *raw performance* sections). Of note when performance level increased (i.e., Top25% and 170 Top10%), RAE magnitude also, expectedly, increased.

When re-analysing distributions following CAPS application, results generally identified a 171 172 removal (or reduction) of RAEs (see Table 2 - corrected performance sections). Specifically, for male 173 jumpers, RAEs predominantly dissipated in both the Top25% and Top10% of performance in all age 174 groups, except 12 years of age. At 12 years of age, CAPs generated an RAE reversal (i.e., disproportionate number of jumpers born in the second-half of the year compared to the first-half) with 175 176 a medium and large effect size in the Top25% and Top10%, respectively. For females, CAPs led to 177 more consistent evenly distributed quartile distributions when compared to distributions in raw data (e.g., 12 and 13 years of age). In fact, no significant odds ratio was apparent for an age-group quartile 178 179 comparison (OR range = 0.57-1.02), suggesting consistent RAE removal.

180

#### 181 Discussion

The purpose of this study was to validate the capability of CAPs to remove RAE inequalities in a unique youth athletic context. The study initially estimated the longitudinal relationships between decimal age and performance in female and male long jumpers. Afterward, CAPs were applied to an independent sample of long jumpers. That latter step enabled the testing of whether CAPs could remove or substantially reduce RAE magnitudes.

Similar to previous studies in national<sup>1</sup> or international<sup>21</sup> sprinters and swimmers (ranged aged: 187 10-18 years),<sup>24,30</sup> results identified significant curvilinear relationships between decimal age and long 188 jump performance for males and females. Based on trendlines long jump performances increased from 189 190 4.18 and 4.04 m at 12 years old to 6.59 and 5.33 m at 17 years for male and female jumpers, respectively. 191 Within the same age-group (e.g., 12 years old), the relative difference in performance estimates from being the relatively youngest (11.01 years) to the relatively oldest (i.e., 12.00) progressed from relatively 192 larger to smaller differences (e.g., about 48-5 % in males 12-17 %; see Table 1). In comparison, for 193 194 females annual jump performance difference magnitudes were lower, with progressively lowering percentage differences within younger age-groups. The earlier (1-2 years) reductions in within year
 performance differences, may also reflect the earlier occurrence of maturation and attainment of adult
 stature.<sup>31</sup>

In Part II, consistent significant overall asymmetries in relative age distributions were observed 198 with large-small effect size in males and medium-small effect size in females (see Table 2 'All'). 199 200 Corroborating with previous RAE findings where samples across age-groups, performance levels and sex were examined within a sport context,<sup>2,7,15,17</sup> the proportion of relatively older long jumpers in this 201 study was on average 2.7 (male) and 1.9 times (female) higher than relatively younger jumpers. While 202 203 effect size decreased with age-group, as expected when (artificial) selection or performance levels were introduced (i.e., Top25% and 10% of performances), RAE sizes again increased.<sup>5,7</sup> Relatively older 204 205 jumpers were approximately 4-5 times more likely to be included in the Top25% and Top10% ranking (see OR Q1 v Q4 in Table 2). For female jumpers, such a trend only existed between 12-14 years age, 206 207 with only descriptive inequalities persisting across 15-17 years of age.

208 When CAPs were applied to the Top25% and 10% of performers, relative age inequalities were 209 predominantly removed, irrespective of sex, age-group, and selection criteria (no significant effect size 210 or OR). In other words, no significant differences in quartile or half year distributions were apparent. 211 For example, focusing on the Top10% of male jumpers at 13 years old, raw data identified a Q1 v Q4 212 distribution OR of 6.86 [3.26, 14.30]. Following CAPs, the equivalent distribution OR was 0.86 [0.47, 213 1.58]. The only exception occurred in male jumpers at 12 years old, where following CAPs a significant 214 RAE reversal in favour of the relatively youngest was apparent for both the Top25% and Top10% of 215 performers (see Table 2). In other words, relatively young long jumpers were more likely to be overrepresented.<sup>21,23</sup> However, this finding should be considered carefully as less initial male 216 217 longitudinal data was available to inform regression estimates and raw sample sizes were comparatively 218 lower in this age-group category. In relation to the latter, as larger percentage distribution changes are 219 more likely in smaller samples, a greater sample would likely help identify greater distribution accuracy. These points reinforce the importance of sufficient sampling, and given current data, suggest a potential 220 221 need for age-range restraint in CAPs application until further data is available.

222 Based on a substantial dataset containing longitudinal data tracking long jump performance 223 across multiple years, present findings predominantly validate and provide efficacy to CAPs application within long-jumping, similar to youth athletic contexts.<sup>1,21,24,27</sup> Practically, results demonstrate the 224 225 capability to more equitably evaluable performance in long jumping while accounting for chronological age related inter-individual differences. CAPs application at junior local-national levels in Italian long-226 227 jumping could help minimize (or at least reduce) the occurrence of RAE bias, help prevent advantages 228 based on earlier anthropometric development as opposed to performers being recognised on the basis of technical and biomechanical skill/proficiency. Such procedures may establish greater accuracy in 229 performance evaluation given the developmental stage of a performer in the sporting context,<sup>24,27</sup> 230 particularly valuable considering coach decision-making, selection, and access to athlete developmental 231 programmes (national federation) is predominantly determined by raw performance criteria<sup>13,14,32,33</sup> 232 Relatedly, coaching knowledge and awareness of RAEs, the occurrence of inter-individual 233 developmental differences, and CAPs could help promote modifications to long jump performance 234 235 evaluation for the betterment of young athletes. Nonetheless, it is still important to understand that other factors, such as the influence of social agents (i.e., as parents, coaches or the athletes themselves), may 236 237 still amplify RAEs.<sup>34</sup> Further, other developmental factors (e.g., maturation status) may also need to be considered.<sup>30</sup> Finally, as CAPs is based on adjustment considerate of a reference population, being able 238 239 to still understand factors impacting inter-individual variability, or developmental change, over time 240 remains problematic.25

### 241 Conclusion

To conclude, present results provide validity and efficacy to CAPs as a strategy to remove RAEs in Italian long-jumping specifically, with potential relevance to other track and field events or youth sport contexts. Local – national governing body stakeholders could consider CAPs as a potential datadriven strategy to help account for inter-individual developmental differences in performance, reduce sport drop-out, increase motivation, and help sustain participation across youth developmental age ranges. CAPs may also help improve the accuracy of performance evaluation and long-term talent identification.

249

#### 250 Practical Implications

251	•	Track & field-related sports systems and practitioners need to be aware of the inter-individual
252		developmental differences (e.g., relative age) occurring within youth age groups which likely
253		lead to participation and performance attainment inequalities.

- Study findings validate the application of Corrective Adjustment Procedures (CAPs) to address
   relative age differences in the context of youth Italian long jumping for males and females,
   although with some caution as to the age-range of CAPs application. If sufficient reference data
   is available, CAPs should be implemented at a sex and age-group matched level within long
   jumping. CAPs could practically be utilised within long-jumping by either event organisers
   and/or individual coaches.
- More broadly, CAPs development and application has the potential to help improve youth
   athlete participation experiences (e.g., self-evaluation and motivation), particularly for those
   commonly disadvantaged by fixed annual-age groups and/or dates applied for youth
   competition. CAPs could help improve accuracy in coach evaluation and talent identification.

## 264 References

265	1.	Romann M, Cobley S. Relative Age Effects in Athletic Sprinting and Corrective Adjustments
266		as a Solution for Their Removal. PLoS One. 2015; 10(4):e0122988.
267	2.	Cobley S, Baker J, Wattie N, McKenna J. Annual age-grouping and athlete development: a
268		meta-analytical review of relative age effects in sport. Sports Med. 2009; 39(3):235-256.
269	3.	Wattie N, Cobley S, Baker J. Towards a unified understanding of relative age effects. J Sports
270		Sci. 2008; 26(13):1403-1409.
271	4.	Barnsley RH, Thompson AH, Barnsley PE. Birthdate and performance: The relative age effect.
272		Can J Hist Sport Phys Educ. 1985; 51:23-28.
273	5.	Cobley S, Baker J, Wattie N, McKenna J. Annual age-grouping and athlete development: a
274		meta-analytical review of relative age effects in sport. Sports Med. 2009; 39(3):235-256.
275	6.	Musch J, Grondin S. Unequal Competition as an Impediment to Personal Development: A
276		Review of the Relative Age Effect in Sport. Dev Rev. 2001; 21(2):147-167.
277	7.	Smith KL, Weir PL, Till K, Romann M, Cobley S. Relative Age Effects Across and Within
278		Female Sport Contexts: A Systematic Review and Meta-Analysis. Sports Med. 2018:1-30.
279	8.	Brustio PR, Lupo C, Ungureanu AN, Frati R, Rainoldi A, Boccia G. The relative age effect is
280		larger in Italian soccer top-level youth categories and smaller in Serie A. PLoS One. 2018;
281		13(4):e0196253.
282	9.	Wattie N, Schorer J, Baker J. The relative age effect in sport: A developmental systems model.
283		Sports Med. 2015; 45(1):83-94.
284	10.	Cobley S, Abbott S, Moulds K, Hogan C, Romann M. Re-balancing the Relative Age Effect
285		Scales Meta-analytical Trends, Causes, and Corrective Adjustment Procedures as a Solution,
286		Chapter 11, in Relative Age Effects in Sport International Perspectives. C. J, Dixon JC, Horton
287		S, Chittle L, J. B, ed^eds. New York, Taylor & Francis, 2020.
288	11.	Till K, Baker J. Challenges and [Possible] Solutions to Optimizing Talent Identification and
289		Development in Sport. Front Psychol. 2020; 11(664).
290	12.	Wattie N, Cobley S, Baker J. Towards a unified understanding of relative age effects. J Sports
291		Sci. 2008; 26(13):1403-1409.
292	13.	Boccia G, Cardinale M, Brustio PR. Elite Junior Throwers Unlikely Remain at the Top Level
293		in the Senior Category. Int J Sports Physiol Perform. 2021; 16(9):1281-1287.
294	14.	Boccia G, Cardinale M, Brustio PR. Performance progression of elite jumpers: Early
295		performances do not predict later success. Scand J Med Sci Sports. 2021; 31(1):132-139.
296	15.	Till K, Cobley S, Wattie N, O'Hara J, Cooke C, Chapman C. The prevalence, influential factors
297		and mechanisms of relative age effects in UK Rugby League. Scand J Med Sci Sports. 2010;
298		20(2):320-329.

- Lupo C, Boccia G, Ungureanu AN, Frati R, Marocco R, Brustio PR. The Beginning of Senior
  Career in Team Sport Is Affected by Relative Age Effect. *Front Psychol.* 2019; 10(1465).
- Brustio PR, Kearney PE, Lupo C, et al. Relative Age Influences Performance of World-Class
  Track and Field Athletes Even in the Adulthood. *Front Psychol.* 2019; 10:1395.
- 18. Kearney PE, Hayes PR, Nevill A. Faster, higher, stronger, older: Relative age effects are most
  influential during the youngest age grade of track and field athletics in the United Kingdom. J
  Sports Sci. 2018; 36(20):2282-2288.
- Hollings SC, Hume PA, Hopkins WG. Relative-age effect on competition outcomes at the
   World Youth and World Junior Athletics Championships. *Eur J Sport Sci.* 2014;
   14(sup1):S456-S461.
- 20. Campbell E, Irving R, Poudevigne M, et al. Contextual factors and sporting success: The
  relationship between birth date and place of early development on the progression of Jamaican
  track and field athletes from junior to senior level. *PLoS One*. 2019; 14(12):e0227144.
- Brustio PR, Boccia G. Corrective procedures remove relative age effect from world-class junior
  sprinters. *J Sports Sci.* 2021.
- Smith KL, Weir PL. Late Birthday Benefits. The "Underdog Hypothesis", in *Relative Age Effects in Sport International Perspectives*. C. J, Dixon JC, Horton S, Chittle L, J. B, ed^eds.
   New York, Taylor & Francis, 2020.
- 23. Cobley S, Abbott S, Dogramaci S, et al. Transient Relative Age Effects across annual age groups
  in National level Australian Swimming. *J Sci Med Sport*. 2018; 21(8):839-845.
- 24. Cobley S, Abbott S, Eisenhuth J, Salter J, McGregor D, Romann M. Removing relative age
  effects from youth swimming: The development and testing of corrective adjustment
  procedures. *J Sci Med Sport.* 2019; 22(6):735-740.
- Webdale K, Baker J, Schorer J, Wattie N. Solving sport's 'relative age'problem: A systematic
  review of proposed solutions. *Int Rev Sport Exerc.* 2020; 13(1):187-204.
- Kelly AL, Jackson DT, Taylor JJ, Jeffreys MA, Turnnidge J. "Birthday-Banding" as a Strategy
  to Moderate the Relative Age Effect: A Case Study Into the England Squash Talent Pathway. *Front Sports Act Living.* 2020; 2(17):573890.
- Abbott S, Moulds K, Salter J, Romann M, Edwards L, Cobley S. Testing the application of
   corrective adjustment procedures for removal of relative age effects in female youth swimming.
   *J Sports Sci.* 2020:1-8.
- Atheltic W. Book of Rules. <u>https://www.worldathletics.org/about-iaaf/documents/book-of-</u>
   <u>rules</u>.
- 332 29. Cohen J. Statistical power analysis for the behavioral sciences, Academic press; 2013.
- 33. Abbott S, Hogan C, Castiglioni MT, et al. Maturity-related developmental inequalities in age33. group swimming: The testing of 'Mat-CAPs' for their removal. *J Sci Med Sport.* 2021;
  335 24(4):397-404.

336	31.	Brustio PR, Cardinale M, Lupo C, Varalda M, De Pasquale P, Boccia G. Being a top swimmer
337		during the early career is not a prerequisite for success: A study on sprinter strokes. J Sci Med
338		Sport. 2021; 24(12):1272-1277.
339	32.	Boccia G, Cardinale M, Brustio PR. World-Class Sprinters' Careers: Early Success Does Not
340		Guarantee Success at Adult Age. Int J Sports Physiol Perform. 2020; 16(3):367-374.
341	33.	Boccia G, Moise P, Franceschi A, et al. Career Performance Trajectories in Track and Field
342		Jumping Events from Youth to Senior Success: The Importance of Learning and Development.
343		PLoS One. 2017; 12(1):e0170744.

344 34. Hancock DJ, Adler AL, Côté J. A proposed theoretical model to explain relative age effects in
345 sport. *Eur J Sport Sci.* 2013; 13(6):630-637.

## 347 Figure Legends

348 Figure 1. Longitudinal quadratic trendline summarizing the relationship between chronological age and

349 raw long jump performance in males (Figure 1 a) and females (Figure 1 b).