



#### AperTO - Archivio Istituzionale Open Access dell'Università di Torino

#### Joint use of cardio-embolic and bleeding risk scorers in elderly patients with atrial fibrillation.

This is the author's manuscript
Original Citation:
Availability:
This version is available http://hdl.handle.net/2318/149221 since
Published version:
DOI:10.1016/j.ejim.2013.08.697
Terms of use:
Open Access
Anyone can freely access the full text of works made available as "Open Access". Works made available under a Creative Commons license can be used according to the terms and conditions of said license. Use of all other works requires consent of the right holder (author or publisher) if not exempted from copyright protection by the applicable law.

(Article begins on next page)



### UNIVERSITÀ DEGLI STUDI DI TORINO

This Accepted Author Manuscript (AAM) is copyrighted and published by Elsevier. It is posted here by agreement between Elsevier and the University of Turin. Changes resulting from the publishing process - such as editing, corrections, structural formatting, and other quality control mechanisms - may not be reflected in this version of the text. The definitive version of the text was subsequently published in

Eur J Intern Med. 2013 Dec;24(8):800-6. doi: 10.1016/j.ejim.2013.08.697. Epub 2013 Sep 12.

You may download, copy and otherwise use the AAM for non-commercial purposes provided that your license is limited by the following restrictions:

(1) You may use this AAM for non-commercial purposes only under the terms of the CC-BY-NC-ND license.

(2) The integrity of the work and identification of the author, copyright owner, and publisher must be preserved in any copy.

(3) You must attribute this AAM in the following format: Creative Commons BY-NC-ND license
(http://creativecommons.org/licenses/by-nc-nd/4.0/deed.en),
+ Digital Object Identifier link to the published journal article on Elsevier's ScienceDirect® platform:
DOI: 10.1016/j.ejim.2013.08.697

# Joint use of cardio-embolic and bleeding risk scores in elderly patients with atrial fibrillation

Maura Marcucci <sup>a, b</sup>, Alessandro Nobili<sup>c</sup>, Mauro Tettamanti<sup>c</sup>, Alfonso Iorio<sup>a, b</sup>, Luca Pasina<sup>c</sup>, Codjo D. Djade<sup>c</sup>, Carlotta Franchi<sup>c</sup>, Alessandra Marengoni<sup>d</sup>, Francesco Salerno<sup>e</sup>, Salvatore Corrao<sup>f</sup>, Francesco Violi<sup>g</sup>, Pier Mannuccio Mannucci<sup>h</sup>,

on behalf of REPOSI Investigators (REPOSI: REgistro POliterapie Società Italiana di Medicina Interna)

a Department of Clinical Epidemiology and Biostatistics, McMaster University, Hamilton, ON, Canada

b Department of Medicine, McMaster University, Hamilton, ON, Canada

cIRCCS-Istituto di Ricerche Farmacologiche "Mario Negri", Milan, Italy

d Geriatric Unit, Ospedali Civili, Department of Medical and Surgery Sciences, University of Brescia, Italy

e Internal Medicine, IRCCS Policlinico San Donato, Department of Medical and Surgery, University of Milan, Italy

f Biomedical Department of Internal Medicine, University of Palermo, Italy

g Divisione I Clinica Medica, Department of Internal Medicine and Medical Specialities, Sapienza University of Rome, Rome, Italy

h Scientific Direction, IRCCS Ca' Granda Maggiore Hospital Foundation, Milan, Italy

Corresponding author at: McMaster University, 1280 Main Street West, Hamilton, ON L8S 4K1, Canada. Tel.: + 1 905 525 9140x20065; fax: + 1 905 526 8447.

#### 1. Introduction

The underuse of vitamin K antagonists (VKAs) among elderly patients with atrial fibrillation (AF) has been confirmed in different settings [1], [2], [3], [4], [5] and [6]. Indeed, the CHA2DS2–VASc [6] and [7], that assigns 2 points (and not 1 as CHADS2[8]) to age  $\geq$  75 years, and 1 point to age  $\geq$  65 years would qualify all patients older than 75 years as candidates for long term anticoagulation, and all patients older than 65 years for aspirin treatment, even in the absence of other risk factors [9], [10] and [11].

Since the fear of treatment-related bleeding is the most likely reason for the underprescription of anticoagulants, tools for the prediction of the risk of bleeding in patients with AF on VKAs have been proposed [12], [13], [14] and [15]. All the available scores for bleeding risk include older age among risk factors. The different therapeutic guidelines frame their recommendations on the degree of cardio-embolic risk based upon CHADS2 or CHA2DS2–VASc, but fail to express uniform agreement on the use and usefulness of bleeding scores, although suggesting of considering the patient bleeding risk to decide on the long-term antithrombotic therapy [9], [10] and [11]. Moreover, it is still controversial whether and to which extent the decisions on cardio-embolic prophylaxis in the most common population of patients with AF (the oldest old with multimorbidity) currently rely upon a joint assessment of both cardio-embolic and bleeding risks.

With this background, we analyze retrospectively patients older than 65 years with atrial fibrillation or flutter (AFF) enrolled in REPOSI [16] during the first (2008) and the second (2010) collection waves of this registry, with the aims to describe: i) the stratification of patients according to the different scores for cardio-embolic and bleeding risks; ii) the within-patient relationship between cardio-embolic and bleeding risks as defined by these scores; iii) whether or not the prescription of antithrombotic therapy was related to the score-based assessment of cardio-embolic and bleeding risks.

#### 2. Materials and methods

#### 2.1. Study population

Patients analyzed in this study were recruited in the frame of the 'REgistro POliterapie SIMI' (REPOSI) [16]. The REPOSI is a collaborative and independent Registry organized by the Italian Society of Internal Medicine (SIMI) and the Mario Negri Institute of Pharmacological Research in Milan with the purpose to create a network of internal medicine and geriatric wards in order to evaluate hospitalized patients older than 65 years affected by multiple diseases and prescribed with polypharmacy. Patients recruited for

REPOSI in 2008 and 2010, and admitted to the participating Italian wards with a known diagnosis of AFF (International Classification of Diseases — Ninth Revision [ICD-9] codes 427.31 or 427.32) were analyzed in this study. Patients newly diagnosed with AFF during the index hospitalization were not included.

#### 2.2. Cardio-embolic and bleeding risk stratification

The patient population was retrospectively classified according to the cardio-embolic risk as predicted by CHADS2 and CHA2DS2–VASc scores [6] and [7], and according to the bleeding risk as predicted by HEMORR2HAGES and HAS-BLED scores [12] and [14]. The components of each score, the annual event rates associated with the risk categories as reported in the literature, as well as the corresponding absolute risk reduction or increase with VKAs are summarized in online Appendix A. The scores were retrospectively calculated for each patient using the data collected at admission on socio-demographic characteristics, clinical history and drug use before the hospitalization and the reason for hospitalization. A modified HEMORR2HAGES score not including genetic risk factors, and a modified HAS-BLED score, not including the labile INR factor were used, because the corresponding data were not available in REPOSI; both these modified versions of the scores have already been used and validated [12], [14] and [17]. The resulting risks were reported both as continuous scores and as categories (low, intermediate, high),

using for the latter the originally proposed score-based stratifications [6], [12], [14], [17] and [18] (online Appendix A). Classification of patients' cardio-embolic risk was compared using both scores, and the classification of patients' bleeding risk using both scores. We then described the co-stratification of the study population using both a scheme for the cardio-embolic risk and one for the bleeding risk, testing different combinations. Correlation between scores, as a measure for trend, was tested by the Spearman test. Concordance/discordance between risk categories was expressed as percentage of patients classified into the same/different risk category. Although risk categories are categorical ordinal variables, linear regression analyses were used to show the average association between the risk categories as defined using the different scores. For this purpose the low, intermediate and high risk categories were coded as 0, 1 and 2, respectively.

#### 2.3. Antithrombotic therapy and risk scores

The study population was characterized according to the antithrombotic therapy recorded at hospital admission, considering as long-term therapy VKAs and antiplatelet drugs (aspirin, clopidogrel, ticlopidine and aspirin plus dypyridamole). To evaluate retrospectively the association between the cardio-embolic/bleeding risk scores and the prescribed antithrombotic therapy, two sets of analyses were performed.

a.

Risk scores as predictors of VKA prescription. A classic logistic regression was used to evaluate this relationship, in simple and multivariable analyses (including both cardioembolic and bleeding score as predictors). CART (Classification and Regression Trees analysis) [19] was also used as a multivariable approach to further explore how the scores were hierarchically associated with VKA prescription. The program automatically selected for each score the best-splitting value for the therapeutic choice, i.e. that value above or below which VKAs were more likely to be prescribed or not.

b.

Risk scores as predictors of antithrombotic therapy type. With the aim of taking into account all the possible antithrombotic options for AFF, a 4-level nominal variable was also used as dependent variable, coded as 0 for no therapy, 1 for antiplatelet therapy, 2 for VKAs, and 3 for VKAs plus antiplatelet agents. The variable levels were chosen in order to simulate an ordinal variable where each further level corresponded to an increasing antithrombotic burden. The association between this variable and the scores was explored using an ordered logistic regression when the proportional odds assumption was met, i.e. when the effect of the score on each therapeutic step was constant (theomodel user's command for STATA was used to verify the assumption). If this assumption was not met, a multinomial logistic regression was used, where the no-therapy choice was taken as reference and the association of the score with any other therapeutic choice was compared to the reference.

Then the analyses exploring the association between the risk scores and antithrombotic therapy were repeated adjusting for patient age, in order to look at the effect of the scores after holding the patient age constant; this is equivalent to remove the effect of age (a component of the scores) from the effect of the scores.

In order to take into account the multi-center origin of the REPOSI data, we adopted robust variance estimates that were obtained in all regression models by means of the Huber/White/sandwich estimator which considers observations as independent across groups (the REPOSI centers in this case).

STATA was used to perform all the analyses (version 12, Statacorp, College Station, Tx, US).

#### 3. Results

#### 3.1. Study population

The 2008–2010 installments of REPOSI included 2712 patients, 1332 enrolled in 2008 and 1380 in 2010. Five hundred and forty-three patients (20.0%) were admitted to hospital with a known diagnosis of AFF, 247 in 2008 (18.5%) and 296 in 2010 (21.4%). Patients with AFF at admission (Table 1) were significantly older than those without (median age = 81.1, range 65.4–100.6 years, versus median age = 78.6, range 65.0–101.4 years, p < 0.001); approximately 80% of patients were older than 75 years. Two hundred sixty-five were males (48.8%), with no difference in gender composition compared to patients without AFF. Twenty-eight patients with AFF at admission (5%) died during the hospitalization. Table 1 shows also the proportion of patients presenting a stroke or a bleeding event as reason for admission or during the hospital stay.

#### 3.2. Cardio-embolic and bleeding risk stratification

Table 1 reports the mean and median score values at admission. Table 2 shows how the study population was stratified into cardio-embolic and bleeding risk categories based upon the different scores. A high correlation was found between the two cardio-embolic risk scores (Spearman correlation coefficient 0.86, p value < 0.001), but with a discordance of 25% between the two risk classifications. In detail, all patients classified at intermediate or high risk using CHADS2 were classified at high risk according to CHA2DS2-VASc; patients with a low cardio-embolic risk according to CHADS2 were reclassified by CHA2DS2-VASc as having an intermediate (9 of 16, 56%) or a high (7 of 16, 43%) risk. There was a high correlation between the two bleeding risk scores (Spearman correlation coefficient 0.82, p value < 0.001), but with a discordance of 43% between the two risk classifications. In detail, nearly all (117 of 119, 98%) patients classified at high risk according to HEMORR2HAGES were classified at high risk also according to HAS-BLED; 57% (193/340) of patients classified at intermediate risk according to HEMORR2HAGES were also classified at intermediate risk according to HAS-BLED, whereas the remaining 43% patients (147/340) were classified at high HAS-BLED risk. Patients with a low bleeding risk according to HEMORR2HAGES were classified at intermediate (83 among 84, 99%) or, in one case only, at high HAS-BLED

risk. Fig. 1 (plots a and b) in the online Appendix B exemplifies the average relationship between risk categories defined by each couple of scores.

Table 3 shows how the study population was co-classified according to both the cardioembolic and bleeding risks using two different score combinations. The Spearman correlation between CHADS2 and HEMORR2HAGES scores and between CHA2DS2– VASc and HAS-BLED scores was, respectively, 0.424 and 0.316. Most of the patients were at high cardio-embolic/high-intermediate bleeding risk (70.5% when CHADS2 plus HEMORR2HAGES were used, 98.3% when CHA2DS2–VASc plus HAS-BLED were used). Plots c, d, e and f in Fig. 1 (online Appendix A) show the average relationship between cardio-embolic and bleeding risk categories using the 4 possible score combinations.

According to the predicted risk associated with the scores reported in the original papers (see the online Appendix A for details): White cells: the predicted annualized cardioembolic risk tends to be larger than the predicted annualized bleeding risk (and the predicted absolute risk reduction of cardio-embolic events with warfarin tends to be larger than the predicted absolute risk increase of bleeding events with warfarin). Dark gray cells: the predicted annualized bleeding risk tends to be larger than the predicted annualized bleeding risk tends to be larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin). Dark gray cells: the predicted absolute risk reduction of cardio-embolic events with warfarin tends to be larger than the predicted absolute risk reduction of cardio-embolic events with warfarin). Light gray cells: the predicted annualized bleeding risk tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted annualized cardio-embolic risk (and the predicted annualized cardio-embolic risk (and the predicted annualized bleeding risk tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted absolute risk reduction of cardio-embolic events with warfarin). Predicted denotes as reported in score validation studies.

#### 3.3. Antithrombotic therapy and risk scores

The antithrombotic therapy that REPOSI patients were receiving at admission is shown in Table 1.

#### 3.3.1. Risk scores as predictors of VKA prescription

Table 4 reports the number and percentage of patients on VKAs in each cell co-defined by the cardio-embolic and bleeding risk. The highest rate of VKA prescription was found among patients at intermediate cardio-embolic and low bleeding risk when the CHADS2/HEMORR2HAGES co-classification was used, and among patients at high

cardio-embolic and intermediate bleeding risk when the CHA2DS2–VASc/HAS-BLED combination was used (ignoring the 100% cell including only 1 patient). In simple logistic regressions, a higher bleeding score, using either HEMORR2HAGES or HAS-BLED, was associated with a lower probability to receive VKA (p < 0.001). Neither cardio-embolic risk score was significantly associated with VKAs prescription in unadjusted analysis. Only after adjusting for the bleeding risk score (either HEMORR2HAGES or HAS-BLED) was a higher cardio-embolic risk score (either CHADS2 or CHA2DS2–VASc) associated with a higher probability to receive VKAs (p < 0.001 for any combination). When all the 4 scores were included as covariates, the HEMORR2HAGES and CHADS2 scores remained significant predictors. The CART analysis confirmed these results, and pointed out that a low bleeding risk score seemed to affect positively the probability of VKA prescription, whereas cardio-embolic risk scores were associated with the probability of VKA prescription.

According to the predicted risk associated with the scores reported in the original papers (see the online Appendix A for details): White cells: the predicted annualized cardioembolic risk tends to be larger than the predicted annualized bleeding risk (and the predicted absolute risk reduction of cardio-embolic events with warfarin tends to be larger than the predicted absolute risk increase of bleeding events with warfarin). Dark gray cells: the predicted annualized bleeding risk tends to be larger than the predicted annualized bleeding risk tends to be larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be larger than the predicted absolute risk reduction of cardio-embolic events with warfarin tends to be larger than the predicted absolute risk reduction of cardio-embolic events with warfarin). Light gray cells: the predicted annualized bleeding risk tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted annualized cardio-embolic risk (and the predicted annualized cardio-embolic risk (and the predicted annualized bleeding risk tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted annualized cardio-embolic risk (and the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted absolute risk increase of bleeding events with warfarin tends to be equal or larger than the predicted annualized cardio-embolic events with warfarin). Predicted absolute risk reduction of cardio-embolic events with warfarin). Predicted denotes as reported in score validation studies.

#### 3.3.2. Risk scores as predictors of the type of antithrombotic therapy

When an ordered 4-level variable was used for antithrombotic therapy, the proportional odds assumption was met for both cardio-embolic risk scores, i.e. higher scores were associated to therapeutic choices with a higher antithrombotic potency, but in a quasi statistically significant way only for CHADS2 (p = 0.054). The proportional odds assumption was not met for the bleeding risk scores. In simple multinomial analysis, and also after adjusting for any cardio-embolic risk score, the HEMORR2HAGES score was

associated with the therapeutic choice, but in different ways: a higher HEMORR2HAGES score was negatively associated with the prescription of VKA compared to no therapy, but it was positively associated with the prescription of antiplatelet agents compared to no therapy. A direct association between a higher score and antiplatelet prescription was also found for the HAS-BLED score in simple multinomial logistic regression, and after adjusting for any cardio-embolic risk score. Conversely HAS-BLED was not associated with the prescription of VKA. None of the reported findings changed when patients on LMWH or fondaparinux were excluded from the analyses. After adjusting for patient age, both the cardio-embolic risk scores became significantly associated with the antithrombotic therapy in all types of analysis even without adjusting for the bleeding risk scores. All the remaining results did not change.

#### 4. Discussion

The REPOSI registry was designed in order to collect data on a representative sample of patients admitted to internal medicine wards, increasingly characterized in Italy and elsewhere in Europe by advanced age and multimorbidity. The first aim of these post-hoc analyses was to describe how the available scores for cardio-embolic and bleeding risks would classify patients with AFF in this complex population. We then evaluated whether or not risk assessment according to the scores was related the choice of antithrombotic therapy.

The main novelty of this study was to look, albeit retrospectively, at the co-classification of this elderly population using a combination of scores for both cardio-embolic and bleeding risk, that might theoretically provide the physician with a higher potential for tailoring each individual treatment than using a strategy based only on the cardio-embolic. As expected, the REPOSI population was on average both at high cardio-embolic and bleeding risk (see Table 2a), even though the patients' cardio-embolic risk category tended to be higher than the bleeding risk category. In particular, the percentage of patients belonging to a cardio-embolic risk category higher than the bleeding risk category was more than 60% when CHADS2 plus HEMORR2HAGES were jointly used, and approximately 50% when CHA2DS2–VASc plus HAS-BLED were used. This co-classification would apparently lead to recommend anticoagulation for approximately 50% of REPOSI patients. However, the same definitions for risk category (i.e. low, intermediate or high) for different scores do not correspond to the same annual risk of stroke or bleeding (and so to the same absolute effect of the treatment), as reported in the online Appendix A. In addition, a more

appropriate way of using the score-based predictions of risk to individualize treatment recommendations should take into account also the different weight that a patient might assign to such clinical events, as stroke and bleeding [20], [21] and [22].

Our data confirm the well-known reclassification effect of CHA2DS2–VASc [6], [7] and [19], which moved almost all patients at low and intermediate CHADS2 score to the high risk category. As expected by definition for a  $\geq$  65 year population, none of the REPOSI patients was classified as having a low CHA2DS2–VASc risk [7] and [19], with the implication that according to this score all REPOSI patients with AFF would be treated with anticoagulants.

The cardio-embolic risk stratification of REPOSI patients resembled that recently described in an elderly cohort from the UK General Practice Research Database (GPRD) [5]. However, the REPOSI population had a higher representation of patients at intermediate-high CHADS2 score, presumably because of a higher mean age and different selection criteria (patients at the time of hospital admission, with a likely higher rate of morbidity than those referred to general practitioners).

There was also a high representation of the high risk category for bleeding among REPOSI patients, higher than in the UK cohort [5]. As for the cardio-embolic scores, a reclassification effect with HAS-BLED was observed compared to HEMORR2HAGES. Indeed, none of the REPOSI patients was at low HAS-BLED risk (because of age, none had a 0 score), and in 40% of them HAS-BLED classified patients into a higher risk category than HEMORR2HAGES. This effect was attenuated provided that a HAS-BLED score of 1 was included in the low risk category together with score 0 (as done in other studies [5] and [18]), yet only 11% of patients had a HAS-BLED score of 1. In fact, HAS-BLED was developed in order to provide a therapeutic guideline easier to memorize and includes more practicable risk factors than HEMORR2HAGES [14].

We observed a low overall rate of prescription of VKAs, confirming a previous analysis based on REPOSI[16]. More interesting, the distribution of the percentages of patients treated with VKAs across the cells defined by the scores (Table 3) and the results of the logistic analyses showed that the patient's bleeding risk, but not the cardio-embolic risk alone, predicted the therapeutic choice. These findings on the relationship between the bleeding score and VKA prescription are consistent with those of the UK cohort [5]. In the literature, evidence on the relationship between cardio-embolic scores and VKA prescription in real settings is not uniform [5], [23] and [24]. In the present study, the cardio-embolic risk was a predictor of VKA prescription, only after adjusting for the

bleeding score or patient age. In addition, the association found between a higher bleeding score and antiplatelet therapy clearly confirms the tendency to prescribe aspirin in clinical practice when evidence or perception of a higher risk of bleeding prevents VKA prescription. Irrespective of the cardio-embolic risk, this situation materializes especially in the elderly, even though this behavior is not justified by a safer profile of aspirin compared to VKAs [25] and [26], and either by a clear efficacy of aspirin [27].

This study has several limitations. First, a certain degree of underreporting is expected because of the post-hoc nature of our research question. Thus, it is possible that the actual risk scores were underestimated. Second, this was only an indirect and theoretical investigation of the association between patients' risks and physician's decisions, because it is not known whether or not REPOSI physicians applied these scores to take decisions. Another limitation is the assumption that the risk scores proposed in the literature for patients with AFF have a good predictive ability in a REPOSI-like elderly population. A further fundamental step should be the evaluation of the impact on patient outcomes of a decision strategy based on combined cardio-embolic and bleeding risk assessment compared to a strategy of cardio-embolic risk assessment alone.

#### Learning points

- Scores for cardio-embolic and bleeding risk in patients with atrial fibrillation are described in the literature to aid at tailoring the long term antithrombotic therapy; all of them include age as risk factor.
- We observed how the available scores (CHADS2 and CHA2DS2–VASc, for cardioembolic risk, and HEMORR2HAGES and HAS-BLED, for bleeding risk) coclassified complex elderly patients with multimorbidity admitted to Italian internal medicine and geriatric wards, and we confirmed that they configured a population both at high cardio-embolic and at high bleeding risk.
- 50–60% of patients (depending on the score couples used) were classified in a cardio-embolic risk category higher than the bleeding risk category.
- In those patients, the prescription and the type of antithrombotic therapy appeared to be primarily influenced by the bleeding risk; both the cardio-embolic scores were associated with the therapeutic choice only after adjusting for the patient bleeding score or age.

#### Conflict of interest

The authors declare that they do not have any conflict of interests.

#### Acknowledgments

We thank Prof. John C. Sinclair at the Department of Clinical Epidemiology and Biostatistics, McMaster University, for the intellectual stimulus, the support and the help provided during the draft of the manuscript.

## Investigators and co-authors of the REPOSI (REgistro POliterapie SIMI, Società Italiana di Medicina Interna) Study Group are as follows

**Steering Committee**: Pier Mannuccio Mannucci (Chair, Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano), Alessandro Nobili (co-chair, Istituto di Ricerche Farmacologiche "Mario Negri", Milano), Mauro Tettamanti, Luca Pasina, Carlotta Franchi (Istituto di Ricerche Farmacologiche "Mario Negri", Milano), Francesco Salerno (IRCCS Policlinico San Donato Milanese, Milano), Salvatore Corrao (ARNAS Civico, Di Cristina, Benfratelli, DiBiMIS, Università di Palermo, Palermo), Alessandra Marengoni (Spedali Civili di Brescia, Brescia), Alfonso Iorio (McMaster University, Hamilton, Canada), Maura Marcucci (McMaster University, Hamilton, Canada).

**Clinical data monitoring and revision**: Valentina Spirito, Damia Noce, Jacopo Bonazzi, Rossana Lombardo, Eleonora Sparacio, Stefania Alborghetti (Istituto di Ricerche Farmacologiche "Mario Negri", Milano).

**Database Management and Statistics:** Mauro Tettamanti, Luigi De Vittorio, Codjo Djignefa Djade (Istituto di Ricerche Farmacologiche "Mario Negri", Milano).

**Investigators:** Domenico Prisco, Elena Silvestri, Caterina Cenci, Tommaso Barnini (Azienda Ospedaliero Universitaria Careggi Firenze, SOD Patologia Medica); Giuseppe Delitala, Stefano Carta, Sebastiana Atzori (Azienda Mista Ospedaliera Universitaria, Sassari, Clinica Medica); Gianfranco Guarnieri, Michela Zanetti, Annalisa Spalluti (Azienda Ospedaliera Universitaria Ospedali Riuniti di Trieste, Trieste, Clinica Medica Generale e Terapia Medica); Maria Grazia Serra, Maria Antonietta Bleve (Azienda Ospedaliera "Cardinale Panico" di Tricase, Lecce, Unità Operativa Complessa Medicina); Massimo Vanoli, Giulia Grignani, Gianluca Casella (Azienda Ospedaliera della Provincia di Lecco, Ospedale di Merate, Lecco, Medicina Interna); Laura Gasbarrone (Azienda Ospedaliera Ospedaliera Ospedaliera Seria); Laura Gasbarrone (Azienda Ospedaliera Ospedaliera Ospedaliera Medica); Massimo Vanoli, Giulia Forlanini, Roma, Medicina Interna 1); Giorgio Maniscalco, Massimo

Gunelli, Daniela Tirotta (Azienda Ospedaliera Ospedale San Salvatore, Pesaro, Soc Medicina Interna); Antonio Brucato, Silvia Ghidoni, Paola Di Corato (Azienda Ospedaliera Papa Giovanni XXIII, Bergamo, Medicina 1); Mauro Bernardi, Silvia Li Bassi, Luca Santi (Azienda Ospedaliera Policlinico Sant'Orsola-Malpighi, Bologna, Semeiotica Medica Bernardi); Giancarlo Agnelli, Alfonso Iorio, Maura Marcucci, Emanuela Marchesini (Azienda Ospedaliera Santa Maria della Misericordia, Perugia, Medicina Interna e Cardiovascolare); Elmo Mannarino, Graziana Lupattelli, Pamela Rondelli, Francesco Paciullo (Azienda Ospedaliera Santa Maria della Misericordia, Perugia, Medicina Interna, Angiologia, Malattie da Arteriosclerosi); Fabrizio Fabris, Michela Carlon, Francesca Turatto (Azienda Ospedaliera Università di Padova, Padova, Clinica Medica I); Maria Cristina Baroni, Marianna Zardo (Azienda Ospedaliera Università di Parma, Parma, Clinica e Terapia Medica); Roberto Manfredini, Christian Molino, Marco Pala, Fabio Fabbian (Azienda Ospedaliera - Universitaria Sant'Anna, Ferrara, Unità Operativa Clinica Medica); Ranuccio Nuti, Roberto Valenti, Martina Ruvio, Silvia Cappelli (Azienda Ospedaliera Università Senese, Siena, Medicina Interna I); Giuseppe Paolisso, Maria Rosaria Rizzo, Maria Teresa Laieta (Azienda Ospedaliera Universitaria della Seconda Università degli Studi di Napoli, Napoli, VI Divisione di Medicina Interna e Malattie Nutrizionali dell'Invecchiamento); Teresa Salvatore, Ferdinando Carlo Sasso (Azienda Ospedaliera Universitaria della Seconda Università degli Studi di Napoli, Napoli, Medicina Interna e Malattie Epato-Bilio Metaboliche Avanzate); Riccardo Utili, Emanuele Durante Mangoni, Daniela Pinto (Azienda Ospedaliera Universitaria della Seconda Università degli Studi di Napoli, Napoli, Medicina Infettivologica e dei trapianti); Oliviero Olivieri, Anna Maria Stanzial (Azienda Ospedaliera Universitaria Integrata di Verona, Verona, Unità Operativa di Medicina Interna B); Renato Fellin, Stefano Volpato, Sioulis Fotini (Azienda Ospedaliera Universitaria Ospedale Sant'Anna, Ferrara, Unità Operativa di Medicina Interna Gerontologia e Geriatria); Mario Barbagallo, Ligia Dominguez, Lidia Plances, Daniela D'Angelo (Azienda Ospedaliera Universitaria Policlinico Giaccone Policlinico di Palermo, Palermo, Unità Operativa di Geriatria e Lungodegenza); Giovanbattista Rini, Pasquale Mansueto, Ilenia Pepe (Azienda Ospedaliera Universitaria Policlinico P. Giaccone di Palermo, Palermo, Medicina Interna e Malattie Metaboliche); Giuseppe Licata, Luigi Calvo, Maria Valenti (Azienda Ospedaliera Universitaria Policlinico P. Giaccone di Palermo, Palermo, Medicina Interna e Cardioangiologia); Claudio Borghi, Enrico Strocchi, Elisa Rebecca Rinaldi (Azienda Ospedaliera Universitaria Policlinico S. Orsola-Malpighi, Bologna, Unità Operativa di Medicina Interna Borghi); Marco Zoli, Elisa Fabbri, Donatella Magalotti (Azienda Ospedaliera Universitaria Policlinico S. Orsola-Malpighi, Bologna, Unità Operativa di Medicina Interna Zoli); Alberto Auteri, Anna Laura Pasqui, Luca Puccetti (Azienda Ospedaliera Universitaria Senese, Siena, Medicina 3); Franco Laghi Pasini, Pier Leopoldo Capecchi, Maurizio Bicchi (Azienda Ospedaliera Universitaria Senese, Siena, Unità Operativa Complessa Medicina 2); Carlo Sabbà, Francesco Saverio Vella, Alessandro Marseglia, Chiara Valentina Luglio (Azienda Ospedaliero-Universitaria Consorziale Policlinico di Bari, Bari, Medicina Interna Universitaria C. Frugoni); Giuseppe Palasciano, Maria Ester Modeo, Annamaria Aquilino, Pallante Raffaele (Azienda Ospedaliero-Universitaria Consorziale Policlinico di Bari, Bari, Medicina Interna Ospedale "Pende-Ferrannini"); Stefania Pugliese, Caterina Capobianco (Azienda Ospedaliero-Universitaria Consorziale Policlinico di Bari, Bari, Clinica Medica I Augusto Murri); Alfredo Postiglione, Maria Rosaria Barbella, Francesco De Stefano (Azienda Ospedaliera Universitaria Policlinico Federico II di Napoli, Medicina Geriatrica Dipartimento di Clinica Medica); Luigi Fenoglio, Chiara Brignone, Christian Bracco, Alessia Giraudo (Azienda Sanitaria Ospedaliera Santa Croce e Carle di Cuneo, Cuneo, S. C. Medicina Interna); Giuseppe Musca, Olga Cuccurullo (Azienda Sanitaria Provinciale di Cosenza Presidio Ospedaliero di Cetraro, Cosenza, Unità Operativa Complessa Medicina Interna); Luigi Cricco, Alessandra Fiorentini (COB Stabilimento Montefiascone, Viterbo, Unità Operativa Complessa di Geriatria e Medicina); Maria Domenica Cappellini, Giovanna Fabio, Sonia Seghezzi, Margherita Migone De Amicis (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Unità Operativa Medicina Interna IA); Silvia Fargion, Paola Bonara, Mara Bulgheroni, Rosa Lombardi (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Medicina Interna 1B); Fabio Magrini, Ferdinando Massari, Tatiana Tonella (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Unità Operativa Medicina Cardiovascolare); Flora Peyvandi, Alberto Tedeschi, Raffaella Rossio (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Medicina Interna 2); Guido Moreo, Barbara Ferrari, Luisa Roncari (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Medicina Interna 3); Valter Monzani, Valeria Savojardo, Christian Folli, Maria Magnini (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Medicina d'Urgenza); Daniela Mari, Paolo Dionigi Rossi, Sarah Damanti, Silvia Prolo (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Geriatria); Maria Sole Lilleri (Fondazione IRCCS Cà Granda Ospedale Maggiore Policlinico, Milano, Medicina Generale ad Indirizzo Geriatrico); Luigi Cricco, Alessandra Fiorentini (COB Viterbo, Stabilimento Montefiascone, Viterbo, UOC Geriatria e Medicina); Giuliana Micale (IRCCS Istituto Auxologico Italiano, Milano, Medicina Generale ad indirizzo Geriatrico); Mauro Podda, Carlo Selmi, Francesca Meda (IRCCS Istituto Clinico Humanitas, Milano, Clinica Medica); Francesco Salerno, Silvia Accordino, Alessio Conca, Valentina Monti (IRCCS Policlinico San Donato e Università di Milano, San Donato Milanese, Medicina Interna); Gino Roberto Corazza, Emanuela Miceli, Marco Vincenzo Lenti, Donatella Padula (IRCCS Policlinico San Matteo di Pavia, Pavia, Clinica Medica I, Reparto 11); Carlo L. Balduini, Giampiera Bertolino, Stella Provini, Federica Quaglia (IRCCS Policlinico San Matteo di Pavia, Pavia, Clinica Medica III); Giovanni Murialdo, Marta Bovio (IRCS Azienda Ospedaliera Universitaria San Martino-IST di Genova, Genova, Clinica di Medicina Interna 2); Franco Dallegri, Luciano Ottonello, Alessandra Quercioli, Alessandra Barreca (Università di Genova, Genova, Medicina Interna 1); Maria Beatrice Secchi, Davide Ghelfi (Ospedale Bassini di Cinisello Balsamo, Milano, Divisione Medicina); Wu Sheng Chin, Laura Carassale, Silvia Caporotundo (Ospedale Bassini, Cinisello Balsamo, Milano, Unità Operativa di Geriatria); Luigi Anastasio, Lucia Sofia, Maria Carbone (Ospedale Civile Jazzolino di Vibo Valentia, Vibo Valentia, Medicina interna); Giancarlo Traisci, Lucrezia De Feudis, Silvia Di Carlo (Ospedale Civile Santo Spirito di Pescara, Pescara, Medicina Interna 2); Giovanni Davì, Maria Teresa Guagnano, Simona Sestili (Ospedale Clinicizzato SS. Annunziata, Chieti, Clinica Medica); Elisabetta Bergami, Emanuela Rizzioli (Ospedale del Delta, Lagosanto, Ferrara, Medicina Interna); Carlo Cagnoni, Luca Bertone, Antonio Manucra (Ospedale di Bobbio, Piacenza, Unità Operativa Medicina e Primo Soccorso); Alberto Buratti, Tiziana Tognin, Nicola Lucio Liberato (Azienda Ospedaliera della Provincia di Pavia, Ospedale di Casorate Primo, Pavia, Medicina Interna); Giordano Bernasconi, Barbara Nardo (Ospedale di Circolo di Busto Arsizio, Varese, Medicina I); Giovanni Battista Bianchi, Sabrina Giaquinto Ospedale "SS Gerosa e Capitanio" di Lovere, Bergamo, Unità Operativa Complessa di Medicina Generale, Azienda Ospedaliera "Bolognini" di Seriate, Bergamo; Giampiero Benetti, Michela Quagliolo, Giuseppe Riccardo Centenaro (Ospedale di Melegnano, Vizzolo Predabissi, Melegnano, Medicina 1); Francesco Purrello, Antonino Di Pino, Salvatore Piro (Ospedale Garibaldi Nesima, Catania, Unità Operativa Complessa di Medicina Interna); Gerardo Mancuso, Daniela Calipari, Mosè Bartone, Francesco Gullo (Ospedale Giovanni Paolo II Lamezia Terme, Catanzaro, Unità Operativa Complessa Medicina Interna); Michele Cortellaro, Marina Magenta, Francesca Perego; Maria Rachele Meroni (Ospedale Luigi Sacco, Milano, Medicina 3°); Marco Cicardi, Antonio Gidaro Marina Magenta (Ospedale Luigi Sacco, Milano, Medicina II); Andrea Sacco, Antonio Bonelli, Gaetano Dentamaro (Ospedale Madonna delle Grazie, Matera, Medicina); Renzo Rozzini, Lina Falanga, Alessandro Giordano (Ospedale Poliambulanza, Brescia, Medicina Interna e Geriatria); Paolo Cavallo Perin, Bartolomeo Lorenzati, Gabriella Gruden, Graziella Bruno (Dipartimento di Scienze Mediche, Università di Torino, Città della Scienza e della Salute, Torino, Medicina 3); Giuseppe Montrucchio, Elisabetta Greco, Pietro Tizzani (Dipartimento di Scienze Mediche, Università di Torino, Città della Scienza e della Salute, Torino, Medicina Interna 5); Giacomo Fera, Maria Loreta Di Luca, Donatella Renna (Ospedale San Giacomo di Monopoli, Bari, Unità Operativa Medicina Interna); Antonio Perciccante, Alessia Coralli (Ospedale San Giovanni-Decollato-Andisilla, Civita Castellana Medicina); Rodolfo Tassara, Deborah Melis, Lara Rebella (Ospedale San Paolo, Savona, Medicina I); Giorgio Menardo, Stefania Bottone, Elsa Sferrazzo (Ospedale San Paolo, Savona, Medicina Interna e Gastroenterologia); Claudio Ferri, Rinaldo Striuli, Rosa Scipioni (Ospedale San Salvatore, L'Aquila, Medicina Interna Universitaria); Raffaella Salmi, Piergiorgio Gaudenzi, Susanna Gamberini, Franco Ricci (Azienda Ospedaliera-Universitaria S. Anna, Ferrara, Unità Operativa di Medicina Ospedaliera II); Cosimo Morabito, Roberto Fava (Ospedale Scillesi d'America, Scilla Medicina); Andrea Semplicini, Lucia Gottardo (Ospedale SS. Giovanni e Paolo, Venezia, Medicina Interna 1); Giuseppe Delitala, Stefano Carta, Sebastiana Atzori (Ospedale Universitario Policlinico di Sassari, Sassari, Clinica Medica); Gianluigi Vendemiale, Gaetano Serviddio, Roberta Forlano (Ospedali Riuniti di Foggia, Foggia, Medicina Interna Universitaria); Luigi Bolondi, Leonardo Rasciti, Ilaria Serio (Policlinico Sant'Orsola-Malpighi, Bologna, Unità Operativa Complessa Medicina Interna); Cesare Masala, Antonio Mammarella, Valeria Raparelli (Policlinico Umberto I, Roma, Medicina Interna D); Filippo Rossi Fanelli, Massimo Delfino, Antonio Amoroso (Policlinico Umberto I, Roma, Medicina Interna H); Francesco Violi, Stefania Basili, Ludovica Perri (Policlinico Umberto I, Roma, Prima Clinica Medica); Pietro Serra, Vincenzo Fontana, Marco Falcone (Policlinico Umberto I, Roma, Terza Clinica Medica); Raffaele Landolfi, Antonio Grieco, Antonella Gallo (Policlinico Universitario A. Gemelli, Roma, Clinica Medica); Giuseppe Zuccalà, Francesco Franceschi, Guido De Marco, Cordischi Chiara, Sabbatini Marta (Policlinico Universitario A. Gemelli, Roma, Roma, Unità Operativa Complessa Medicina d'Urgenza e Pronto Soccorso); Martino Bellusci, Donatella Setti, Filippo Pedrazzoli (Presidio Ospedaliero Alto Garda e Ledro, Ospedale di Arco, Trento, Unità Operativa di Medicina Interna Urgenza/Emergenza); Giuseppe Romanelli, Caterina Pirali, Claudia Amolini (Spedali Civili di Brescia, Brescia, Geriatria); Enrico Agabiti Rosei, Damiano Rizzoni, Luana Castoldi (Spedali Civili di Brescia, Brescia, Seconda Medicina); Antonio Picardi, Umberto Vespasiani Gentilucci, Chiara Mazzarelli, Paolo Gallo (Università Campus Bio-Medico, Roma, Medicina Clinica-Epatologia); Luigina Guasti, Luana Castiglioni, Andrea Maresca, Alessandro Squizzato, Sara Contini, Marta Molaro (Università degli Studi dell'Insubria, Ospedale di Circolo e Fondazione Macchi, Varese, Medicina Interna I); Giorgio Annoni, Maurizio Corsi Sara Zazzetta (Università degli studi di Milano-Bicocca Ospedale S. Gerardo, Monza, Unità Operativa di Geriatria; Marco Bertolotti, Chiara Mussi Roberto Scotto, Maria Alice Ferri, Francesca Veltri (Università di Modena e Reggio Emilia, AUSL di Modena, Modena, Nuovo Ospedale Civile, Unità Operativa di Geriatria); Franco Arturi, Elena Succurro, Giorgio Sesti, Umberto Gualtieri (Università degli Studi Magna Grecia, Policlinico Mater Domini, Catanzaro, Unità Operativa Complessa di Medicina Interna); Francesco Perticone, Angela Sciacqua, Michele Quero, Chiara Bagnato (Università Magna Grecia Policlinico Mater Domini, Catanzaro, Unità Operativa Malattie Cardiovascolari Geriatriche); Paola Loria, Maria Angela Becchi, Gianfranco Martucci, Alessandra Fantuzzi, Mauro Maurantonio (Università di Modena e Reggio Emilia, Medicina Metabolica-NOCSAE, Baggiovara, Modena); Roberto Corinaldesi, Roberto De Giorgio, Mauro Serra, Valentina Grasso, Eugenio Ruggeri, Lorenzo Mauro Carozza, Fabio Pignatti (Dipartimento di Scienze Mediche e Chirurgiche, Unità Operativa di Medicina Interna, Università degli Studi di Bologna/Azienda Ospedaliero-Universitaria S. Orsola-Malpighi, Bologna).

#### Table 1.

#### Demographic and clinical characteristics of the study population<sup>a</sup>.

Male, n (%)	265 (48.8)
Mean age ± SD (median, range)	81.0 ± 7.3 (81.1, 65.4– 100.6)
Median number of drugs per patient (range)	6 (1–15)
Mean CHADS <sub>2</sub> ± SD (median, range)	2.2 ± 1.1 (2, 0–6)
Mean CHA <sub>2</sub> DS <sub>2</sub> -VASc ± SD (median, range)	3.8 ± 1.2 (4, 1–9)
Mean HEMORR <sub>2</sub> HAGES ± SD (median, range)	2.6 ± 1.2 (3, 0–7)
Mean HAS-BLED ± SD (median, range)	2.6 ± 1.1 (2, 1–6)
Oral antithrombotic therapy at admission, n (%) <sup>b</sup>	

Characteristic

Vitamin K antagonist	210 (38.7)
Antiplatelet agent	174 (32.0)
VKA + antiplatelet	16 (3.0)
None	143 (26.3)
Stroke as reason for admission, n (%)	22 (4.0)
Stroke as adverse event during the hospital stay, n (%)	2 (0.4)
Bleeding as reason for admission, n (%)	16 (2.9) <sup>c</sup>
Bleeding as adverse event during the hospital stay, n (%)	8 (1.5) <sup>c</sup>

<sup>a</sup> The risk scores were calculated counting the risk factors at admission.

<sup>b</sup> 15% of patients not receiving VKAs at admission were on low molecular weight heparin (LMWH) or fondaparinux at therapeutic or prophylactic doses.

c Two patients presented a bleeding event both as reason for admission and during the hospital stay.

#### Table 2.

Risk stratification according to cardio-embolic and bleeding scores.

a. Cardio-embolic risk scores

Cardio-embolic risk category	CHADS	<b>D</b> <sub>2</sub>		CHA <sub>2</sub> DS <sub>2</sub> -VASc			
	Score	Number of patients	%	Score	Number of patients	%	
Low	0	16	3.0	0	0	0	
Intermediate	1	118	21.7	1	9	1.7	
High	≥ 2	409	75.3	≥ 2	534	98.3	
b. Bleeding risk scores							
Bleeding risk category	HEMOR	RR₂HAGES		HAS-BLED			
	Score	Number of patients	%	Score	Number of patients	%	
Low	0–1	84	15.5	0	0	0	
Intermediate	2–3	340	62.6	1–2 <sup>A</sup>	278	51.2	
High	≥ 4	119	21.9	≥ 3	265	48.8	

A 60 patients (11.0%) had a HAS-BLED score 1.

#### Table 3.

Patient distribution according to cardio-embolic and bleeding risk categories: number of patients (% of the whole population).

HEMORR <sub>2</sub> HAGES CHADS <sub>2</sub>	Low risk	Intermediate risk	High risk	Total
Low risk	11 (2.0)	5 (0.9)	¥	
Intermediate risk	47 (8.7)	62 (11.4)	9(1.7)	
High risk	26 (4.8)	273 (50.3)	110 (20.2)	
	1. (1.2) II.1.			543 (100)
HAS-BLED CHA2DS2-VASc	Low risk	Intermediate risk	High risk	Total
Low risk	4	2	1	
Intermediate risk	-	8(1.5)	1 (0.2)	
High risk	<u> </u>	270 (49.7)	264 (48.6)	

#### Table 4.

Frequency of VKA prescription according to cardio-embolic and bleeding risk categories: number of patients (% of the total number of patients in each cell).

HEMORR <sub>2</sub> HAGES	Low risk	Intermediate risk	High risk	
Low risk	4 (36.4)	1 (20.0)	4	5 (31.2)
Intermediate risk	31 (66.0)	19 (30.6)	0 (0.0)	40 (33.9)
High risk	16 (61.5)	130 (47.6)	25 (22.7)	171 (41.8)

HAS-BLED CHA2DS2-VASc	Low risk	Intermediate risk	High risk	
Low risk	+	-	-	(e)
Intermediate risk	-	2 (25.0)	1 (100.0)	3 (33.3)
High risk	ā	150 (55.6)	73 (27.6)	223 (41.8)

appendix A: Cardio-embolic and bleeding risk scores.

appendix B: Figures.

#### Supplementary material

Online appendix A. Cardio-embolic and bleeding risk scores

#### Box 1. Stroke Risk Stratification with CHADS<sub>2</sub> and CHA<sub>2</sub>DS<sub>2</sub>-VASc score

CHADS₂ items	Points
Congestive heart failure/left ventricular dysfunction	1
Hypertension	1
Aged ≥ 75 years	1
Diabetes mellitus	1
Stroke/TIA/systemic embolism	2
Maximum score 6	
CHA <sub>2</sub> DS <sub>2</sub> -VASc items	Points
Congestive heart failure/left ventricular dysfunction	1
Hypertension	1
Aged ≥ 75 years	2
Diabetes mellitus	1
Stroke/TIA/systemic embolism	2
Vascular disease (prior MI, PAD or aortic plaque)	1
Sex gender	1
Age 65-74 years	1
Maximum cooro 9	

Legend: TIA, Transient Ischemic Attack. MI, Myocardial Infarction. PAD, Peripheral Artery Disease

#### CHADS<sub>2</sub> risk categories

Low risk: score 0 Intermediate risk: score 1 High risk: score ≥ 2 CHA<sub>2</sub>DS<sub>2</sub>-VASc risk categories Low risk: score 0

Intermediate risk: score 1 High risk: score ≥ 2 Box 2. Bleeding risk stratification with  $HEMORR_2HAGES$  and HAS-BLED score

HEMORR <sub>2</sub> HAGES items	Points
Hepatic or renal disease	1
Ethanol abuse	1
Malignancy	1
Aged ≥ 75 years	1
Reduced platelet count or function*	1
Rebleeding risk	2
Hypertension	1
Anemia	1
Genetic factors	1
Excessive fall risk or neuropsychiatric disease	1
Stroke	1
Maximum score 12	
HAS-BLED items	Points
Hypertension	1
Abnormal renal or liver function	1 each
Stroke	1
Bleeding	1
Labile INR	1
Elderly (aged ≥ 65 years)	1
Drugs <sup>∫</sup> or alcohol	1 each
Maximum score 9	

\*Aspirin use. <sup>(</sup>Antiplatelet agents or nonsteroidal anti-inflammatory drugs.

#### HEMORR<sub>2</sub>HAGES risk categories

Low risk: score 0-1 Intermediate risk: score 2-3 High risk: score ≥ 4 HAS-BLED risk categories Low risk: score 0 Intermediate risk: score 1-2 High risk: score ≥ 3

Вох	3.	Score-based	risk	categories	and	reported	annualized	event	rate	as	events/100	person-years
(cor	fid	ence interval	in th	e original st	udy) <sup>;</sup>	*						

	Low Risk	Intermediate Risk	High Risk
CHADS <sub>2</sub> [8]	1.9 (1.2-3.0)	2.8 (2.0-3.8)	4.0-18.2 (3.1 – 27.4)
CHA <sub>2</sub> DS <sub>2</sub> –VASc [6]	0	0.7	1.9-14.2
CHA <sub>2</sub> DS <sub>2</sub> –VASc [22]	0.6	1.3	2.8-15.9
HEMORR <sub>2</sub> HAGES [12]	1.9-2.5 (0.6-4.3)	5.3-8.4 (3.4-13.6)	10.4-12.3 (5.1-23.1)
HEMORR <sub>2</sub> HAGES [18]	3.1 (2.8-3.3)	6.3 (5.9-6.7]	12.2 (11.1-13.3)
HAS-BLED [22]	2.8	7.1-12.4	14.9-35.4
HAS-BLED [18]	2.7 (2.4-2.9) <sup>∫</sup>	5.5 (5.2-6.0)	8.1 (7.6-8.6)

\*For cardio-embolic events among patients off VKA; for bleeding events, among patients on VKA <sup>#</sup>Back-calculated from the rates in patients off VKA, considering a RR of 2.3 for bleeding with warfarin <sup>J</sup>HAS-BLED scores 0 and 1 were included in low risk category

Box 4. Score-based risk categories and associated absolute risk reduction of cardio-embolic events (per 100 person-years) and absolute risk increase of bleeding events (per 100 person-years) with warfarin, using a relative risk for cardio-embolic events of 0.36 and a relative risk for bleeding event of 2.30 with warfarin vs placebo

	Low Risk	Intermediate Risk	High Risk
CHADS <sub>2</sub> [8]	1.2	1.8	2.6-11.6
CHA <sub>2</sub> DS <sub>2</sub> –VASc [6]	0	0.4	1.2-9.1
CHA <sub>2</sub> DS <sub>2</sub> –VASc [22]	0.4	0.8	1.8-10.2
HEMORR <sub>2</sub> HAGES [12]	1.1-1.4	3.0-4.7	5.9-6.9
HEMORR <sub>2</sub> HAGEs [18]	1.7	3.6	6.9
HAS-BLED [22]	1.6	4.0-7.0	8.4-20.0
HAS-BLED [18]	1.5 <sup>∫</sup>	3.1	4.6

\*For cardio-embolic events among patients off VKA; for bleeding events, among patients on VKA <sup>#</sup>Back-calculated from the rates in patients off VKA, considering a RR of 2.3 for bleeding with warfarin <sup>J</sup>HAS-BLED scores 0 and 1 were included in low risk category Online appendix B.

Figures.

Figure legend.

### Figure 1. Relationship between risk categories according to cardio-embolic and bleeding risk scores, based on standard linear regressions

The six plots show the relationship between risk categories as defined by couples of scores. Assuming that the classification into risk categories can approximate an ordinal variable, a value of 0, 1 and 2 was assigned to, respectively, a low, intermediate and high risk category. The x- and y-axis labels reflect this coding: 0 =low risk category; 1 = intermediate risk category; 2 = high risk category. In each plot the results for the following regression parameters are provided: b = regression coefficient, a = intercept (constant), p value for test for statistical significance of the regression coefficient. The relationship between each couple of scores can be summarized as: risk category based on score y = a + b(risk category based on score x), and result should be interpreted accounting for the categorical nature of the variables. The more the line resembles the 45° line (a=0 and b=1), the closer the co-classification is to the perfect coincidence between cardio-embolic and bleeding risk category.

The fist two plots (a and b)at the top of the figure depict the relationship between cardio-embolic risk categories *or* bleeding risk categories as defined by different scores. For example patients at low CHADS<sub>2</sub> score ("0") belonged on average to a CHA<sub>2</sub>DS<sub>2</sub>–VASc category of "1.79", i.e. some patients were classified at intermediate ("1"), some patients at high CHA<sub>2</sub>DS<sub>2</sub>–VASc risk ("2"); patients at intermediate CHADS<sub>2</sub> score (1) belonged on average to a CHA<sub>2</sub>DS<sub>2</sub>–VASc category of "1.9" (=1.79+1\*0.11), i.e. practically all the patients were reclassified as at high CHA<sub>2</sub>DS<sub>2</sub>–VASc risk ("2"); the flat line points out the fact the almost all the population was at high CHA<sub>2</sub>DS<sub>2</sub>–VASc score. Similarly, patients at low HEMORR<sub>2</sub>HAGES (0) risk belonged on average to a HAS-BLED category of 0.96 ( $\cong$ 1, i.e. almost all were at intermediate risk according to HASBLED); patients at intermediate HEMORR<sub>2</sub>HAGES risk (category "1") belonged on average to a HAS-BLED.

The plots c, d, e and f, showing the relationship between cardio-embolic risk categories *and* bleeding risk categories using different combination of scores, can be similarly interpreted. Overall the regression equations point out that, independently of the couple of scores used (even if less evidently when CHA<sub>2</sub>DS<sub>2</sub>–VASc and HAS-BLED are compared), the patient cardio-embolic risk category was on average higher than the bleeding risk category.

#### Figure 2. CART analysis: risk scores as predictors of VKA prescription

The figure shows the results of a CART analysis where all the scores were entered as independent variables and the prescription of VKAs (yes/no) was the outcome. In the figure: N = number of patients in each branch; F = number of failures (= patients prescribed with VKA) in each branch; RPR = Relative Prescription Ratio (where a RPR=1 represents a hypothetical point of indifference for VKA prescription) for each branch. The analyses retained the CHADS<sub>2</sub> and the HEMORR<sub>2</sub>HAGES scores as the most significant predictors of VKAs prescription. In particular, the HEMORR<sub>2</sub>HAGES score was that best-splitting the population between those more and less likely to receive VKAs, with a score of 3 as that theoretically overturning the chance of being prescribed with VKAs. The value of CHADS<sub>2</sub> score appeared to have a secondary effect: among patients less likely to receive VKAs because of their higher HEMORR<sub>2</sub>HAGES score, a high value of CHADS<sub>2</sub> score seemed to increase the chance of VKA prescription.



Figure 1. Relationship between risk categories according to cardio-embolic and bleeding risk scores, based on standard linear regressions

#### Figure 2. CART analysis: risk scores as predictors of VKA prescription.

CART analysis - Split if (adjusted) P<.05. Dependent variable: VKA prescription at admission Indep. variables: CHADS<sub>2</sub>, CHA<sub>2</sub>DS<sub>2</sub>-VASc, HEMORR<sub>2</sub>HAGES, HAS-BLED\*



N = number of patients in each branch; F = number of failures (= patients prescribed with VKA) in each branch; RPR = Relative Prescription Ratio